# DS 122 Homework 4

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## Contents

1 Q1 Problem A	3
1.1 Q1.1 Part 1	3
1.2 Answer	3
1.3 Q1.2 Part 2	4
1.4 Answer	4
1.5 Q1.3 Part 3	4
1.6 Answer	5
1.7 Q1.4 Part 4	5
1.8 Answer	5
1.9 Q1.5 Part 5	6
1.10 Answer	6
2 Q2 Problem B	7
2.1 Q2.1	8
2.2 Q2.2	8
2.3 Q2.3	8
3 Q3 Problem C	8
3.1 Q3.1 Part 1	8
3.2 Answer	9
3.3 Q3.2 Part 2	9
3.4 Answer	10
3.5 Q3.3 Part 3	
3.6 Answer	
4 Q4 Problem D	11
4.1 Q4.1	12
4.2 Answer	
4.3 Q4.2	13
4.4 Answer	13
4.5 Q4.3	13
4.6 Answer	13
4.7 Q4.4	14
4.8 Answer	14
5 Q5 Problem E	15
5.1 Answer	15
6 Q6 Computational	16

## 1 Q1 Problem A

## 1.1 Q1.1 Part 1

Differentiate the expression  $y(x) = \log(x^3) + \log(4x) - \log(3x + 2)$ 

## 1.2 Answer

Given:

$$y(x) = \log(x^3) + \log(4x) - \log(3x + 2)$$

For  $\log(x^3)$ 

Let  $u = x^3$ . Then:

$$u' = 3x^2$$

The derivative of  $\log(x^3)$  is:

$$\frac{3x^2}{x^3} = \frac{3}{x}$$

For  $\log(4x)$ 

Let u' = 4x. Then:

$$u'=4$$

Thus, the derivative of log(4x) is:

$$\frac{4}{4x} = \frac{1}{x}$$

For  $\log(3x+2)$ 

Let u = 3x + 2. Then:

$$u'=3$$

Thus, the derivative of  $\log(3x+2)$  is:

$$\frac{3}{3x+2}$$

Combining the results:

The overall derivative, y'(x), is the sum of the derivatives of each item:

$$y'(x) = \frac{3}{x} + \frac{1}{x} - \frac{3}{3x+2}$$

### 1.3 Q1.2 Part 2

Given the functions  $u(x) = e^x$  and  $v(x) = \log(x+1)$ , determine the derivative of y(x) = u(x)v(x)

#### 1.4 Answer

To find the derivative of y(x) = u(x)v(x), we'll need to make use of the product rule for differentiation. The product rule states:

If y(x) = u(x)v(x)

$$y'(x) = u'(x)v(x) + u(x)v'(x)$$

Given:

$$u(x) = e^x$$

$$v(x) = \log(x+1)$$

Differentiating u(x):

$$u'(x) = \frac{d}{dx}e^x = e^x$$

Differentiating v(x):

Using the chain rule:

$$v'(x) = \frac{d}{dx}\log(u) = \frac{u'}{u}$$

Let u = x + 1. Then:

$$u'=1$$

Thus, the derivative of log(x + 1) is:

$$v'(x) = \frac{1}{x+1}$$

Applying the Product Rule

$$y'(x) = u'(x)v(x) + u(x)v'(x)$$

$$y'(x) = e^x \log(x+1) + e^x \frac{1}{x+1}$$

## 1.5 Q1.3 Part 3

Using the logarithmic properties to simplify, find the derivative of  $y(x) = \log(x^4 \cdot (x+1)^3)$ 

### 1.6 Answer

Given:  $y(x) = \log(x^4 \cdot (x+1)^3)$ 

Using the Properties of Logarithms:

Using the logarithmic property  $\log(a \cdot b) = \log(a) + \log(b)$ 

$$y(x) = \log(x^4) + \log((x+1)^3)$$

Using the property  $\log(a^b) = b \cdot \log(a)$ :

$$y(x) = 4\log(x) + 3\log(x+1)$$

Differentiating the Expression:

Differentiating  $4 \log(x)$ 

$$\frac{d}{dx}4\log(x) = 4 \cdot \frac{1}{x} = \frac{4}{x}$$

Differentiating  $3\log(x+1)$ :

Using the chain rule:

$$\frac{d}{dx}3\log(x+1) = 3 \cdot \frac{1}{x+1} = \frac{3}{x+1}$$

Combining the Results:

The overall derivative, y'(x), is the sum of the derivatives of each term:

$$y'(x) = \frac{4}{x} + \frac{3}{x+1}$$

## 1.7 Q1.4 Part 4

For the function  $y(x) = x^2 \log(x^3 + 1)$ , apply the chain rule and logarithmic differentiation to determine  $\frac{dy}{dx}$ .

#### 1.8 Answer

Given the function:

$$y(x) = x^2 \log(x^3 + 1)$$

We can use both the product rule and the chain rule combined with logarithmic differentiation Let's call:

$$u(x) = x^2$$

$$v(x) = \log(x^3 + 1)$$

Using the product rule:

$$\frac{d}{dx} = u'(x)v(x) + u(x)v'(x)$$

Differentiating u(x):

$$u'(x) = \frac{d}{dx}x^2 = 2x$$

Differentiating v(x):

$$v'(x) = \frac{d}{dx} \log(x^3 + 1)$$

$$v'(x) = \frac{1}{x^3 + 1} * \frac{d}{dx}(x^3 + 1)$$

$$v'(x) = \frac{1}{x^3 + 1} * 3x^2$$

$$v'(x) = \frac{3x^2}{x^3 + 1}$$

Applying the product rule:

$$\frac{dy}{dx} = 2x\log(x^3 + 1) + x^2 \frac{3x^2}{x^3 + 1}$$

## 1.9 Q1.5 Part 5

If  $y(x) = x^x$ , take the natural logarithm of both sides and then differentiate implicitly to directly determine  $\frac{dy}{dx}$ .

#### 1.10 Answer

Given  $y(x) = x^x$ 

Taking the Natural Logarithm of Both Sides

$$ln(y) = ln(x^x)$$

Using the property of logarithms  $\ln(a^b) = b \ln(a)$ :

$$\ln(y) = x \ln(x)$$

Differentiating the left side with respect to x using the chain rule:

$$\frac{1}{y} * \frac{dy}{dx}$$

Differentiating the right side:

The term  $x \ln(x)$  is a product of two functions, so we use the product rule.

Let u(x) and  $v(x) = \ln(x)$ 

$$u'(x) = 1$$

$$v'(x) = \frac{1}{x}$$

Using the product rule:

$$\frac{d}{dx}[x\ln(x)] = x * \frac{1}{x} + \ln(x) * 1 = 1 + \ln(x)$$

Equating both derivatives:

$$\frac{1}{u} * \frac{dy}{dx} = 1 + \ln(x)$$

Solving for  $\frac{dy}{dx}$ 

$$\frac{dy}{dx} = y(1 + \ln(x))$$

We know from the original equation that  $y=x^x$ . Plugging this in:

$$\frac{dy}{dx} = x^x (1 + \ln(x))$$

## 2 Q2 Problem B

Consider the histograms below which depict the sampling distributions of four different estimators for a population parameter. The true population parameter value is 7.

Based on the histograms:

### 2.1 Q2.1

Find the estimators with the lowest bias.

- Both 1 and 2 have lowest bias
- Choice 2 of 4: Both 1 and 3 have lowest bias
- Choice 3 of 4: Only 3 has lowest bias
- Choice 4 of 4: Both 2 and 4 have lowest bias

#### 2.2 Q2.2

Find the estimator(s) with the highest variance.

- Choice 1 of 4: Both 1 and 4 have highest variance
- Choice 2 of 4: Both 2 and 4 have highest variance
- Choice 3 of 4: Only 1 has highest variance
- Choice 4 of 4: Only 2 has highest variance

## 2.3 Q2.3

Considering the fact that Mean Squared Error (MSE) is a combination of variance and squared bias, which estimator(s) likely has the highest MSE?

- Choice 1 of 4: Estimator 1
- Choice 2 of 4: Estimator 2
- Choice 3 of 4: Estimator 1 and 2
- Choice 4 of 4: Estimator 4

## 3 Q3 Problem C

Suppose  $\{y^{(1)}, y^{(2)}, ..., y^{(n)}\}$  is an i.i.d. sample from the continuous uniform distribution with parameters 0 and  $\theta$ .

Let 
$$\hat{\theta} = \frac{2}{n} * \left(y^{(1)}, y^{(2)}, ..., y^{(n)}\right)$$
 be an estimator of  $\theta$ 

Hint: You can use the following facts about the continuous uniform distribution with parameters a and b:

- its means is equal to <sup>a+b</sup>/<sub>2</sub>
   its variance is equal to <sup>(b-a)<sup>2</sup></sup>/<sub>12</sub>

## 3.1 Q3.1 Part 1

Determine the bias of  $\theta$ .

#### 3.2 Answer

Given:

 $\{y^{(1)}, y^{(2)}, ..., y^{(n)}\}$  is an i.i.d. sample from the continuous uniform distribution with parameters 0 and  $\theta$ . Estimator:

$$\hat{\theta} = \frac{2}{n} \sum_{i=1}^n y^{(i)}$$

We want to determine the bias of  $\hat{\theta}$ .

Determine  $E(\hat{\theta})$ 

Given  $y^{(i)}$  is from a uniform distribution with parameters 0 and  $\theta$ , the mean  $E\big(y^{(i)}\big)$  is:

$$E\big(y^{(i)}\big) = \frac{0+\theta}{2} = \frac{\theta}{2}$$

Now, the expected value of the sum of the y((i)) values is:

$$E\bigg(\sum_{i=1}^n y((i))\bigg) = nE\big(y^{(i)}\big) = n*\frac{\theta}{2} = \frac{n\theta}{2}$$

Given the estimator:

$$\hat{\theta} = \frac{2}{n} \sum_{i=1}^n y^{(i)}$$

The expected value  $E(\hat{\theta})$  is:

$$E(\hat{\theta}) = \frac{2}{n} * E\left(\sum_{i=1}^{n} y((i))\right)$$

$$E(\hat{\theta}) = \frac{2}{n} * \frac{n\theta}{2} = \theta$$

Compute the Bias

Now that we have  $E(\hat{\theta}) = \theta$ :

$$\operatorname{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta = \theta - \theta = 0$$

Therefore, the estimator  $\hat{\theta}$  is unbiased for  $\theta$ .

## 3.3 Q3.2 Part 2

Determine the mean of  $\widehat{\theta}$ .

### 3.4 Answer

To determine the mean of  $\theta$ , we need to find its expected value,  $E(\hat{\theta})$ .

Recall from the previous part, the estimator is:

$$\hat{\theta} = \frac{2}{n} \sum_{n} ()$$

Given that  $y^{(i)}$  is from a uniform distribution with parameters 0 and  $\theta$ , the mean  $E(y^{(i)})$  is:

$$E\big(y^{(i)}\big) = \frac{0}{\theta}, 2) = \frac{\theta}{2}$$

Now, the expected value of the sum of the  $y^{(i)}$  values is:

$$E(\sum_{i=1}^n y^{(i)} n E\big(y^{(i)}\big) = n \cdot \frac{\theta}{2} = \frac{(n \cdot \theta)}{2}$$

Given the estimator:

$$\hat{\theta} = \frac{2}{n} \sum_{(i=1)}^{n} y^{(i)}$$

The expected value  $E(\hat{\theta} \text{ is:}$ 

$$E(\hat{\theta}) = \frac{2}{n} \cdot E\left(\sum_{i=1}^{n} y^{(i)}\right)$$

$$E(\hat{\theta}) = \frac{2}{n} \cdot \frac{n\theta}{\theta} = \theta$$

Thus, the mean of  $\hat{\theta}$  is  $\theta$ .

## 3.5 Q3.3 Part 3

Determine the variance of  $\hat{\theta}$ .

#### 3.6 Answer

To determine the variance of  $\hat{\theta}$ , we will start by calculating the variance of a single observation  $y^{(i)}$  and then extend it to be the variance of the estimator.

Given:

 $y^{(i)}$  are i.i.d samples from the continuous uniform distribution with parameters 0 and  $\theta$ 

The variance  $\mathrm{Var}ig(y^{(i)}ig)$  is:

$$\operatorname{Var}(y^{(i)}) = \frac{(\theta - 0)^2}{12} = \frac{\theta^2}{12}$$

Since the observations are independent, the variance of the sum is the sum of the variances:

$$\operatorname{Var}\left(\sum_{i=1}^{n} y((i))\right) = n \cdot \frac{\theta^{2}}{12} = \frac{n\theta^{2}}{12}$$

Now, let's consider the estimator:

$$\hat{\theta} = \frac{2}{n} \sum_{i=1}^{n} y(i)$$

Using the properties of variance:

$$Var(cX) = c^2 Var(X)$$

where c is a constant.

The variance  $\operatorname{Var}(\hat{\theta})$  is:

$$\operatorname{Var}(\hat{\theta}) = \left(\frac{2}{n}\right)^2 \operatorname{Var}(X)$$

$$\operatorname{Var}\!\left(\hat{\theta}\right) = \frac{4}{n^2} \cdot \frac{\left(n\theta\right)^2}{12}$$

$$\operatorname{Var}\!\left(\hat{\theta}\right) = \frac{4\theta^2}{12n}$$

$$\operatorname{Var}(\hat{\theta}) = \frac{\theta^2}{3n}$$

Thus, the varaince of  $\hat{\theta}$  is  $\frac{\theta^2}{3n}$ .

## 4 Q4 Problem D

Suppose y is a discrete random variable with the following probability mass function (pmf):

у	0	1	2	3
$p(y;\theta)$	$\frac{2}{3}\theta$	$\frac{1}{3}\theta$	$\frac{2}{3}(1-\theta)$	$\frac{1}{3}(1-\theta)$

Where  $0 \le \theta \le 1$  is a parameter. Given the following 10 independent observations from this distribution:

$$Y = \{1, 1, 0, 2, 2, 1, 3, 2, 0, 3\}$$

#### 4.1 Q4.1

Compute the likelihood function  $p(Y; \theta)$ .

#### 4.2 Answer

We can multiply together the probabilities assigned by the pmf to each observed va;ie. according to the principle of likelihood.

Given observations  $Y = \{1, 1, 0, 2, 2, 1, 3, 2, 0, 3\}$ , let's define:  $n_0, n_1, n_2, n_3$  as the number of times 0, 1, 2, 3 appear in Y respectively.

The likelihood function,  $L(\theta; Y)$ , is computed as:

$$L(\theta; Y) = p(Y; \theta) = p(y_1; \theta)p(y_2; \theta)...p(y_{10}; \theta)$$

Since the observations are independent, we multiply the probabilities. Given the provided pmf:

$$p(y;\theta) =$$

We can compute the likelihood function as:

$$L(\theta;Y) = \left(\frac{2}{3}\theta\right)^{n_0} \left(\frac{1}{3}\theta\right)^{n_1} \left(\frac{2}{3}\theta(1-\theta)\right)^{n_2} \left(\frac{1}{3}\theta(1-\theta)\right)^{n_3}$$

Where:

$$n_0 = \text{number of 0s in Y}$$

$$n_1 = \text{number of 1s in Y}$$

$$n_2 = \text{number of 2s in Y}$$

$$n_3 = \text{number of 3s in Y}$$

Given the observations  $Y = \{1, 1, 0, 2, 2, 1, 3, 2, 0, 3\}$ , we can substitute:

$$n_0 = 2$$

$$n_1 = 3$$

$$n_2 = 3$$

$$n_3 = 2$$

Now, substituting these counts into the likelihood function we can get:

$$L(\theta;Y) = \left(\frac{2}{3}\theta\right)^2 \left(\frac{1}{3}\theta\right)^3 \left(\frac{2}{3}\theta(1-\theta)\right)^3 \left(\frac{1}{3}\theta(1-\theta)\right)^2$$

#### 4.3 Q4.2

Compute the log-likelihood function.

#### 4.4 Answer

The log-likelihood function is the natural logarithm of the likelihood function. Given our likelihood function from Q4.1:

$$L(\theta;Y) = \left(\frac{2}{3}\theta\right)^2 \left(\frac{1}{3}\theta\right)^3 \left(\frac{2}{3}\theta(1-\theta)\right)^3 \left(\frac{1}{3}\theta(1-\theta)\right)^2$$

We can take the natural logarithm of both sides:

$$l(\theta; Y) = \ln[L(\theta; Y)]$$

Applying the properties of logarithms:

$$l(\theta;Y) = 2\ln\left(\frac{2}{3}\theta\right) + 3\ln\left(\frac{1}{3}\theta\right) + 3\ln\left(\frac{2}{3}(1-\theta)\right) = 2\ln\left(\frac{1}{3}(1-\theta)\right)$$

Break it down further:

$$l(\theta;Y) = 2\bigg[\ln\bigg(\frac{2}{3}\bigg) + \ln(\theta)\bigg] + 3\bigg[\ln\bigg(\frac{1}{3}\bigg) + \ln(\theta)\bigg] + 3\bigg[\ln\bigg(\frac{2}{3}\bigg) + \ln(1-\theta)\bigg] + 2\bigg[\ln\bigg(\frac{1}{3}\bigg) + \ln(1-\theta)\bigg]$$

$$l(\theta;Y) = 2\ln\left(\frac{2}{3}\right) + 2\ln(\theta) + 3\ln\left(\frac{1}{3}\right) + 3\ln(\theta) + 3\ln\left(\frac{2}{3}\right) + 3\ln(1-\theta) + 2\ln\left(\frac{1}{3}\right) + 2\ln(1-\theta)$$

#### 4.5 Q4.3

Find the potential Maximum Likelihood Estimator (MLE) for  $\theta$ .

#### 4.6 Answer

We will need to maximize the log-likelihood function in the respect to  $\theta$  to find the Maximum Likelihood Estimator (MLE) for  $\theta$ .

From Q4.2, our log-likelihood function is:

$$l(\theta;Y) = 2\ln\left(\frac{2}{3}\right) + 2\ln(\theta) + 3\ln\left(\frac{1}{3}\right) + 3\ln(\theta) + 3\ln\left(\frac{2}{3}\right) + 3\ln(1-\theta) + 2\ln\left(\frac{1}{3}\right) + 2\ln(1-\theta)$$

To find the MLE, we will differentiate  $l(\theta; Y)$  with respect to  $\theta$  and set the result equal to zero:

$$\frac{dl(\theta; Y)}{d\theta} = 0$$

Differentiating:

- 1. Derivative of  $\ln(\theta)$  is  $\frac{1}{\theta}$
- 2. Derivative of  $\ln(1-\theta)$  is  $-\frac{1}{1-\theta}$
- 3. Constants and terms that don't include  $\theta$  will have derivatives of zero

Setting all this to zero:

$$\frac{5}{\theta} - \frac{5}{1 - \theta} = 0$$
$$5(1 - \theta) - 5\theta = 0$$
$$\theta = \frac{1}{2}$$

So, the Maximum Likelihood Estimator (MLE) for  $\theta$  is  $\theta = 0.5$ .

## 4.7 Q4.4

Explain in words how to confirm that the computed point for part c is indeed a maximum. Note: no computations are required for this part.

#### 4.8 Answer

- 1. Second Derivative Test
  - By computing the second derivative of the log-likelihood function with respect to  $\theta$  and evaluating it at the computed point, we can determine the nature of the point:
    - If the second derivative is negative, it indicates that the function is concave down at that point, which means our computed point is a local maximum.
    - If the second derivative is positive, it indicates that the function is concave up, which would mean our computed point is a local minimum.
    - If the second derivative is zero, the test is inconclusive.
- 2. Log-Likelihood Surface
  - If we were to plot the log-likelihood function, the computed point would be a peak on this surface if it is a maximum.
  - We cam look for the highest point in this plot.
- 3. Sign Change of First Derivative
  - If the first derivative changes sign (from positive to negative) as θ crosses the computed value, this
    indicates that the function increases up to our point and then decreases after, suggesting it's a maximum.

## 5 Q5 Problem E

In a clinical trial for a new drug, 30 patients were treated. Out of them, 24 patients showed improvement in their conditions.

Assuming the patients' responses follow a binomial distribution, find the Maximum Likelihood Estimator (MLE) for the probability  $\theta=p$  that a randomly chosen patient shows improvement after being treated with the drug.

Hint: The likelihood function for binomial distribution is given by:

$$p(x;\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$$

where:

- n is the number of trials
- x is the number of successes
- $\theta$  is the parameter, that represents the probability of success on a single trial.

#### 5.1 Answer

The goal here is to find the value of  $\theta$  (the probability of a patient showing improvement) that maximizes the likelihood of observing the given data, which follows a binomial distribution.

Given:

- n = 30 (total number of patients)
- x = 24 (number of patients who showed improvement)

The likelihood function for the binomial distribution is:

$$L(\theta) = p(x; \theta) = \binom{n}{x} \theta^x (1 - \theta)^{n-x}$$

The Maximum Likelihood Estimator (MLE) of  $\theta$  is the value of  $\theta$  that maximizes this likelihood function.

To find the MLE, we'll differentiate the log likelihood with respect to  $\theta$  and set it to zero. This is because the log function is a strictly increasing function, so where the log likelihood achieves its maximum, the likelihood also achieves its maximum.

Taking the natural logarithm of both sides:

$$\begin{split} \ln(L(\theta)) &= \ln\biggl(\binom{n}{x}\theta^x(1-\theta)^{n-x}\biggr) \\ \ln(L(\theta)) &= \ln\biggl(\binom{n}{x}\biggr) + x\ln(\theta) + (n-x)\ln(1-\theta) \end{split}$$

Now, differentiate with respect to  $\theta$ :

$$\frac{d}{d\theta}\ln(L(\theta)) = \frac{x}{\theta} - \frac{n-x}{1-\theta}$$

To find the MLE, set the derivative to zero and solve for  $\theta$ :

$$\frac{x}{\theta} - \frac{n-x}{1-\theta} = 0$$

Multiplying through by  $\theta(1-\theta)$  and rearranging:

$$x(1-\theta) = (n-x)\theta$$
$$x - x\theta = n\theta - x\theta$$
$$x = n\theta$$

Thus, the MLE for  $\theta$  is:

$$\hat{\theta} = \frac{x}{n}$$

Given x = 24 and n = 30:

$$\hat{\theta} = \frac{24}{30} = \frac{4}{5} = 0.8$$

Therefore, the Maximum Likelihood Estimator (MLE) for the probability  $\theta$  that a randomly chosen patient shows improvement after being treated with the drug is 0.8 or 80%.

## 6 Q6 Computational

See the Jupyter Notebook for details.